

# Modeling mood variation associated with smoking: an application of a heterogeneous mixed-effects model for analysis of ecological momentary assessment (EMA) data

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## ABSTRACT

**Aims** Mixed models are used increasingly for analysis of ecological momentary assessment (EMA) data. The variance parameters of the random effects, which indicate the degree of heterogeneity in the population of subjects, are considered usually to be homogeneous across subjects. Modeling these variances can shed light on interesting hypotheses in substance abuse research. **Design** We describe how these variances can be modeled in terms of covariates to examine the covariate effects on between-subjects variation, focusing on positive and negative mood and the degree to which these moods change as a function of smoking. **Setting** The data are drawn from an EMA study of adolescent smoking. **Participants** Participants were 234 adolescents, either in 9th or 10th grades, who provided EMA mood reports from both random prompts and following smoking events. **Measurements** We focused on two mood outcomes: measures of the subject's negative and positive affect and several covariates: gender, grade, negative mood regulation and smoking level. **Findings and conclusions** Following smoking, adolescents experienced higher positive affect and lower negative affect than they did at random, non-smoking times. Our analyses also indicated an increased consistency of subjective mood responses as smoking experience increased and a diminishing of mood change.

**Keywords** Adolescent smoking, complex variation, diary methods, EMA, experience sampling, heteroscedasticity, log-linear variance, multi-level, variance modeling.

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## INTRODUCTION

Linear mixed models (LMMs, also referred to as multi-level or hierarchical linear models) have become a primary method for analysis of clustered data [1,2], where a cluster may be a subject observed on several occasions or a group of similar subjects. For data clustered within subjects, a basic characteristic of these models is the inclusion of random subject effects into linear models in order to account for the similarity among repeated observations taken from individual subjects. These random effects consist of unmeasured variables reflecting each person's influence on his/her data, and the variance of these random effects indicate the degree of between-person variation (i.e. heterogeneity) in the

population of subjects. Typically, the variance of the random effects is treated as homogeneous across groups of subjects or levels of covariates. However, these homogeneity of variance assumptions can be relaxed by modeling differences in variances across subject groups or attributable to subject covariates.

The study of individual variability has received increasing attention [3–6]; these papers describe many of the conceptual issues and some traditional statistical approaches for examining such variation. LMMs can be used to systematize and extend this work by assessing the determinants of inter-individual or between-subjects variation.

Modeling variances requires a fair amount of data. Modern data collection procedures, such as ecological

momentary assessments (EMA) [7,8], experience sampling [9,10] and diary methods [11], provide this opportunity. These procedures yield relatively large numbers of subjects and observations per subject, and data from such designs are sometimes referred to as intensive longitudinal data [12]. Such designs are in keeping with the 'bursts of measurement' approach described by Nesselrode & McCollam [13], who called for such an approach in order to assess individual variability. As they note, such bursts of measurement increase the research burden in several ways; however, they are necessary for studying individual variation.

Mixed-model analysis of EMA data has been well described by Schwartz & Stone [14]. Additionally, Moghaddam & Ferguson [15] analyzed EMA data using mixed models to examine smoking-related changes in mood. These papers focus on the effects of covariates, either subject-varying or time-varying, on the EMA mean responses. Here we extend this approach by examining the degree to which covariates influence the variation inherent in the EMA data. In this regard, a few articles have described approaches for examining determinants of between- and within-subjects variance from EMA studies. Penner *et al.* [16] used basic descriptive statistical methods to examine relationships among within-subjects variation in several mood variables. More recently, Hedeker *et al.* [17] and Hedeker & Mermelstein [18] have described mixed-model approaches incorporating a log-linear structure for determinants of the within-subjects variance. In this paper, we extend this by modeling the between-subjects variance, allowing covariates to potentially influence the variances associated with the random subject effects. In particular, we focus on the variation of mood that is associated with smoking, and the degree to which subject characteristics influence the mood variation. To aid in making this class of models accessible to researchers, we provide sample computer syntax and output at <http://www.uic.edu/~hedeker/long.html>.

### Adolescent smoking, mood and variability

Many prominent models of cigarette smoking maintain that smoking is reinforcing, and that smoking can relieve negative affect [19,20]. Indeed, both adults and adolescents often claim that smoking is relaxing and reduces emotional distress [21,22]. However, although the relationship between mood and smoking has received substantial empirical attention for adult smokers, much less is known about the acute changes in mood with smoking among adolescents. The present study, with its focus on real-time assessments of mood and smoking among adolescents, helps to shed light on this important topic.

Although there is substantial consensus among both smokers and researchers that smoking helps to regulate

affect, most of the empirical work investigating the smoking–mood relationship has focused on the examination of changes in mean levels of mood with smoking. Surprisingly, although affect regulation inherently implies the modulation of variability in mood as well, the examination of variability in mood and smoking has been largely neglected. As Hertzog & Nesselrode [4] note, describing mean levels of variables is not always adequate for examining key features of developmental change. Variation also conveys important information about the phenomenon of interest. In the case of adolescent smoking and the development of dependence, variation in mood and mood changes may help to explain more clearly the development of tolerance. Examining individual variability may enhance our ability to predict changes in smoking behavior above and beyond what can be achieved by examining mean information alone.

Important, too, in the examination of mood and smoking, is the distinction between within-person and between-person variability. Kassel and colleagues [19,22] have argued persuasively for the need to differentiate causal, within-person mechanisms from between-person data. Whether smoking relieves negative affect is essentially a within-person question, and thus analytical models need to similarly differentiate between within-subject and between-subject effects.

Much of the research on mood and smoking has also been limited to assessments of negative affect, while ignoring positive affect. This neglect is particularly problematic given the theoretical importance of differentiating between negative reinforcement models of smoking and positive reinforcement models, especially in the development of dependence among adolescents [23]. There is also considerable evidence to support the notion that positive and negative affect are distinct constructs, and not just opposite ends of a continuum [24,25]. Thus, in the current study, we assessed both positive and negative affect.

Finally, there may well be individual differences in the extent to which adolescents' moods vary, and whether they vary with smoking. Identifying potential moderator variables may also help in the prediction of smoking escalation among relatively novice smokers. Thus, the aims addressed in this study are: (1) to examine the variation in mood that is associated with smoking and (2) to examine the degree to which subject characteristics influence the mood variation. We hypothesized that a key moderator of mood variability would be an individual's level of expectancies about their own ability to regulate negative moods—negative mood regulation (NMR) [26]. Individuals who hold high expectancies about their ability to cope with negative affect or stressors may show less variability in their mood states. In addition, we hypothesized that level of smoking would also affect

mood variability and changes with smoking. Following along the lines of the development of tolerance with dependence, we hypothesized that as smoking level or experience increased, mood responses to smoking would decrease, as would variability in overall mood.

## METHODS

### Subjects

The data for this paper come from a longitudinal study of the natural history of smoking among adolescents. The study uses a multi-method approach to assess adolescents at multiple time-points (baseline, 6, 9, 15, 24 and 33 months). The data collection modalities include paper and pencil questionnaires and in-person interviews, and for subsets of participants, more intensive measurement modalities including family observations, psychophysiological assessments and week-long time/event EMA sampling via hand-held palmtop computers (referred to as 'electronic diary'). We report here on the data from the baseline EMA collection.

The design of the electronic diary study involved sampling 9th and 10th graders at baseline who had tried smoking at least once during the past 12 months, but who had not yet progressed to smoking five or more cigarettes a day; 461 adolescents completed the baseline assessment for this study. The majority (57.6%) had smoked at least one cigarette in the past month at baseline. Adolescents in the electronic diary study were recruited as part of the larger study of Social-Emotional Contexts of Adolescent Smoking Patterns (total  $n = 1263$ ). Active, written parental consent and adolescent assent were required for participation in the study.

Data collection occurred via hand-held palmtop computers, programmed specifically for our data collection needs, with all other residing programs disabled. Each data collection wave included 7 consecutive days of monitoring. Four types of interviews were programmed into the electronic diary: random prompts, and three types of smoking-related event recordings. Random time prompts were initiated by the device approximately five times per day. Each random prompt was date- and time-stamped and recorded whether the interview was completed, missed, delayed or disbanding. Compliance with responding to the random prompts was very good: approximately 71% of the random prompts were completed. No participants were excluded because of problems in using the devices. The random interviews asked about mood, activity, location, companionship (with whom or alone), presence of other smokers and other behaviors. In addition to the random prompts, participants were trained to event record smoking episodes, as well as episodes when either they had had the opportu-

nity to smoke, but decided actively not to smoke, or when they wanted to smoke, but did not have the opportunity to do so. The 'smoke' and 'non-smoking' interviews included the same questions as the random prompts, and in addition asked about specific smoking-related items (e.g. how much smoked, how the cigarette was obtained, etc.). Random prompts and the self-initiated smoking records were mutually exclusive; no smoking occurred during random prompts.

Because of our interest in comparing mood from random prompts and smoking events, we included in the analysis only subjects who provided data from at least one smoking event during the EMA study phase. In all, there were 234 such subjects with data from a total of 8179 random prompts and smoking events. The average number of random prompts was approximately 30 per subject (median = 30, range = 7–71), and the average number of smoking events was about five per subject (median = 3, range = 1–42). The Spearman correlation between the number of random prompts and number of smoking events was near-zero ( $-0.08$ ) and not statistically significant. This analysis sample of 234 participants included 54.3% female ( $n = 127$ ) and 47.4% 9th graders ( $n = 111$ ). Their ethnic distribution was 59.8% white ( $n = 140$ ), 21.8% Hispanic ( $n = 51$ ), 12.8% African American ( $n = 30$ ) and 5.6% ( $n = 13$ ) other. This subsample was not statistically different from those excluded in terms of gender, grade or race. They also did not differ on psychosocial measures of perceived stress, life events, depressive symptomatology or self-reported grades in school. However, as expected because of inclusion criteria, the adolescents who were included in the analyses smoked significantly more than those excluded; at baseline, adolescents who were included smoked an average of 1.3 cigarettes/day [standard deviation (SD) = 2.36] in the past 7 days, whereas those excluded smoked an average of 0.03 cigarettes/day (SD = 0.12;  $t = -8.36$ ,  $df = 455$ ,  $P < 0.0001$ ). The adolescents included in the analyses, compared to those excluded, also had higher scores on a measure of alcohol problems (assessing frequency and amount of drinking and consequences) at baseline (mean = 4.6, SD = 1.50 compared to mean = 3.5, SD = 1.49,  $t = -7.48$ ,  $df = 459$ ,  $P < 0.001$ ).

### Measures

#### *Negative and positive affect*

Two mood outcomes were considered: measures of the subject's negative and positive affect (denoted NA and PA, respectively) at each random prompt and at each smoking episode. Both these measures consisted of the average of several individual mood items, each rated from 1 to 10, that were identified via factor analysis. Specifically, PA consisted of the following items that reflected

subjects' assessments of their positive mood just before the prompt signal: I felt happy, I felt relaxed, I felt cheerful, I felt confident, and I felt accepted by others. Similarly, NA consisted of the following items assessing pre-prompt negative mood: I felt sad, I felt stressed, I felt angry, I felt frustrated, and I felt irritable. Subjects rated each item on a 1–10 Likert-type scale, with '10' representing very high levels of the attribute. For the smoking events, participants rated their mood immediately after smoking. Over all prompts and events, both random and smoking, and ignoring the clustering of the data within subjects, the mean of PA was 6.77 (SD = 1.96), while the NA mean was 3.53 (SD = 2.28).

#### Gender and grade

To illustrate our approach, we selected a limited number of covariates. First, we considered gender and grade-level with the variables Male (coded 0 = female or 1 = male) and Grade 10 (coded 0 = 9th or 1 = 10th grade).

#### Negative mood regulation expectancies

We also examined a measure of negative mood regulation (NMR) as a covariate because we hypothesized individuals' expectancies about their abilities to manage negative moods would be related to both their mean mood and level of variation. Negative mood regulation expectancies were assessed through the NMR scale developed by Catanzaro & Mearns [26], which was designed to measure the extent to which individuals believe they can do something to alleviate their own negative mood states. This 30-item measure asked the adolescents to indicate how much they agree or disagree (on a five-point Likert scale) with statements beginning with the stem, 'When I'm upset, I believe that . . .'. On this uni-factorial scale, higher scores indicate stronger beliefs about one's ability to regulate negative mood. Scores on NMR ranged from 1 to 5, with higher values indicating more negative mood regulation, and the sample mean was 3.5 (SD = 0.71).

#### Smoking level

Finally, as a measure of a subject's smoking level, we used the number of cigarettes smoked per day in the last 30 days (denoted as SmkLevel). Based on the frequencies, we recoded this variable into six levels: 0 = did not smoke ( $n = 41$ ), 1 = less than one cigarette per day ( $n = 37$ ), 2 = one cigarette per day ( $n = 41$ ), 3 = two cigarettes per day ( $n = 48$ ), 4 = three to five cigarettes per day ( $n = 51$ ) and 5 = more than five cigarettes per day ( $n = 16$ ).

All covariates were measured at baseline and prior to the collection of the EMA data. This helps to explain why there were 41 subjects who, at baseline, indicated that

they did not smoke in the last 30 days (SmkLevel = 0), but who did provide at least one smoking event during the EMA phase (our criterion for being included in the analysis data set). Given the low and infrequent levels of smoking reported by this sample, and the selection criteria excluding more regular smokers, it is not surprising to see continued sporadic patterns of smoking between our baseline assessment week and the EMA assessment week. Of the covariates, there were only two significant positive correlations among them: Male with NMR (Spearman's  $r = 0.26$ ,  $P < 0.0001$ ), and Grade 10 with SmkLevel (Spearman's  $r = 0.15$ ,  $P < 0.02$ ).

#### Data analysis

##### Model 1: standard LMM for changes in the level of mood associated with smoking events

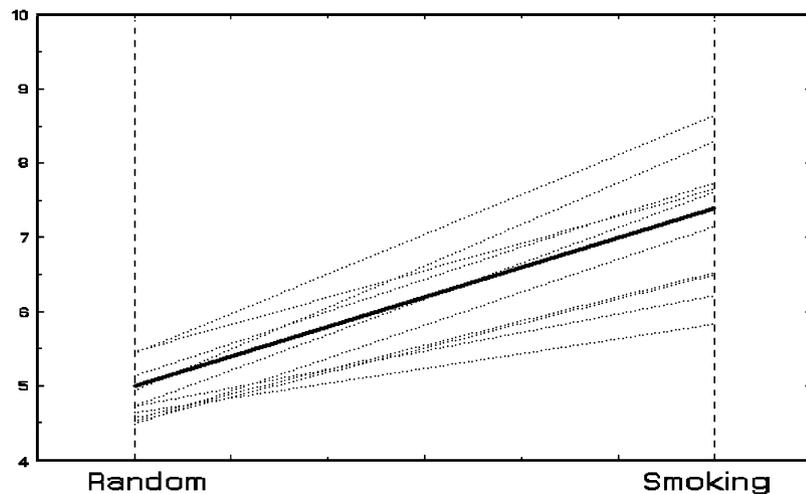
Consider the following linear mixed model (LMM) for the mood measurement  $y$  of individual  $i$  ( $i = 1, 2, \dots, n$  subjects) at occasion  $j$  ( $j = 1, 2, \dots, n_i$  prompts and events), where SmkEvent represents a variable indicating whether the occasion is from a random prompt (= 0) or a smoking event (= 1):

$$y_{ij} = (\beta_0 + \nu_{0i}) + (\beta_1 + \nu_{1i})\text{SmkEvent}_{ij} + \epsilon_{ij} \quad (\text{eqn 1})$$

Hereafter, we use the term and variable 'SmkEvent' to refer to this indicator variable which contrasts smoking events relative to random prompts. This model represents the regression of the outcome variable  $y$  on the independent variable SmkEvent, where  $\beta_0$  is the overall population intercept,  $\beta_1$  is the overall population slope,  $\nu_{0i}$  is the intercept deviation for subject  $i$ ,  $\nu_{1i}$  is the slope deviation for subject  $i$  and  $\epsilon_{ij}$  is an independent error term distributed normally with mean 0 and variance  $\sigma_\epsilon^2$ . The errors are independent conditional on both  $\nu_{0i}$  and  $\nu_{1i}$ . With two random subject-specific effects, the population distribution of intercept and slope deviations is assumed to be a bivariate normal  $N(0, \Sigma_\nu)$ , where  $\Sigma_\nu$  is the  $2 \times 2$  variance-covariance matrix given as:

$$\Sigma_\nu = \begin{bmatrix} \sigma_{\nu_0}^2 & \sigma_{\nu_0\nu_1} \\ \sigma_{\nu_0\nu_1} & \sigma_{\nu_1}^2 \end{bmatrix}$$

This model indicates the effect of SmkEvent both at the individual ( $\nu_{0i}$  and  $\nu_{1i}$ ) and population ( $\beta_0$  and  $\beta_1$ ) levels. Specifically, the intercept parameters indicate the levels of mood during background random prompts, and the slope parameters indicate the degree of mood change associated with smoking events compared to random. The population intercept and slope parameters represent the random mood and mood difference associated with smoking for the population of subjects, whereas the individual parameters express how the individual deviates from the population in terms of their mood.



**Figure 1** Random intercept and slope model

For a visual representation of the model consider Fig. 1, which illustrates the average change in mood attributable to smoking (the bold line) and trend lines (change in mood between random and smoking) of 10 individual subjects who deviate randomly relative to the average trend. The average line is determined by  $\beta_0$  (average mood for random prompts) and  $\beta_1$  (average change in mood for smoking events). Similarly, the individual lines are determined by  $v_{0i}$  (i.e. how different a person is relative to the random prompt average) and  $v_{1i}$  (how different a person is compared to the average slope). The degree of individual mood variation for the random prompts is characterized by  $\sigma_{v_0}^2$ , and the degree of individual mood variation in the slopes (or mood changes for smoking, relative to random) is given by  $\sigma_{v_1}^2$ . Note that in Fig. 1, there is a fair amount of individual variation for both: the 10 individual lines do not equal the bold line in terms of the intercept or slope.

To model more fully the effect of SmkEvent on mood, as described in Begg & Parides [27], we also included the subject's mean  $\overline{\text{SmkEvent}}_i$  as a covariate. Note that because  $\text{SmkEvent}_{ij}$  is simply a binary variable, taking on values of 0 or 1,  $\overline{\text{SmkEvent}}_i$  simply equals the proportion of occasions (i.e. both random prompts and smoking events) that were smoking events for a subject. The model is now written as:

$$y_{ij} = (\beta_0 + v_{0i}) + (\beta_1 + v_{1i})\text{SmkEvent}_{ij} + \beta_2 \overline{\text{SmkEvent}}_i + \varepsilon_{ij} \tag{eqn 2}$$

Here,  $\beta_2$  represents the subject-level effect of  $\overline{\text{SmkEvent}}$ , namely, the association of a person's proportion of smoking events with their average mood across both smoking events and random prompts. Conversely,  $\beta_1$  is the within-subjects effect of SmkEvent, which indicates how a person's mood differs between a random prompt and smoking event, controlling for the proportion of smoking events that the person has. The subject-specific

effects  $v_{0i}$  and  $v_{1i}$  indicate how subjects deviate from these overall effects. That is,  $v_{0i}$  represents a subject's deviation in mood, adjusted for their proportion of smoking events, and  $v_{1i}$  is a subject's deviation in the within-subject SmkEvent effect on mood (also adjusted for their proportion of smoking events) or, in other words, how a given subject's reported mood differs following smoking a cigarette, relative to a random prompt.

Covariates can be added to the model to account for other determinants of mood. For example, to control for any mood-related changes attributable to day of week, one could include six indicator variables in the model. Let  $x_{ij}$  denote a vector of such covariates associated with subject  $i$  and occasion  $j$ . In general, this vector can include both time-invariant and time-varying covariates, as well as their interactions. The model is given as

$$y_{ij} = (\beta_0 + v_{0i}) + (\beta_1 + v_{1i})\text{SmkEvent}_{ij} + \beta_2 \overline{\text{SmkEvent}}_i + \beta'_x x_{ij} + \varepsilon_{ij}, \tag{eqn 3}$$

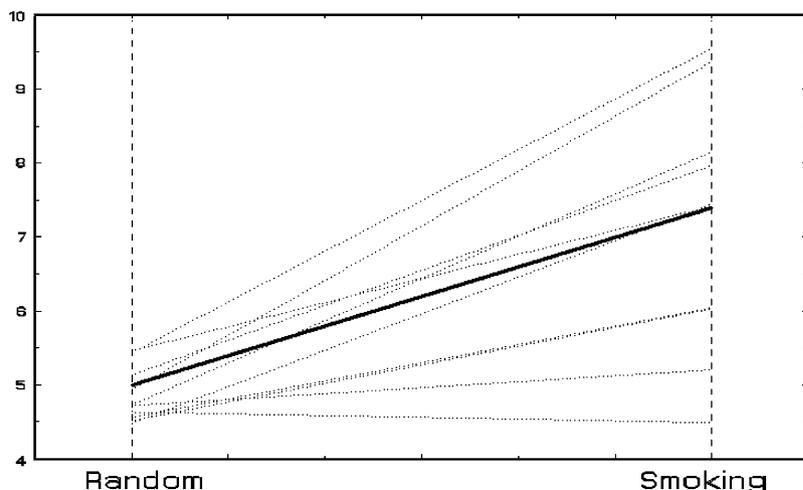
where  $\beta_x$  is the (column) vector of coefficients associated with the additional covariates.

To be consistent with the generalization that will be described below (as model II), let us re-express the random-effect variances associated with the intercepts  $v_{0i}$  and SmkEvent slopes  $v_{1i}$  as:

$$\sigma_{v_0}^2 = \exp(\alpha_{00}), \tag{eqn 4}$$

$$\sigma_{v_1}^2 = \exp(\alpha_{10}), \tag{eqn 5}$$

The reason for using the exponential function will become clear and be described in model II below. For now,  $\alpha_{00}$  and  $\alpha_{10}$  simply represent the intercept and slope variance, respectively, on the natural log scale (because, taking logs,  $\ln \sigma_{v_0}^2 = \alpha_{00}$  and  $\ln \sigma_{v_1}^2 = \alpha_{10}$ ). Specifically,  $\alpha_{00}$  represents mood variation during random prompts (i.e. when  $\text{SmkEvent}_{ij} = 0$ ) in natural log units and  $\alpha_{10}$  represents the variation, in natural log units, of the mood



**Figure 2** Random intercept and slope model with increased slope variance

changes associated with smoking events (i.e. when  $\text{SmkEvent}_{ij} = 1$ ). Because the intercepts  $v_{0i}$  and slopes  $v_{1i}$  are allowed to be correlated (where  $\sigma_{v_0v_1}$  is the covariance), the total mood variation during smoking events is given by  $\exp(\alpha_{00}) + \exp(\alpha_{10}) + 2\sigma_{v_0v_1}$ . Note that in Fig. 1 the amount of variation for smoking events (horizontal spread around the bold line at Smoking) is greater than for random prompts (horizontal spread around the bold line at Random), and so in terms of the parameters, this figure depicts a situation where  $\exp(\alpha_{10}) + 2\sigma_{v_0v_1} > 0$ .

*Model II: heterogeneous LMM for modeling mood variation associated with smoking events*

The LMM in equation 3, which we refer to as model I, includes two random effects: one for a subject's intercept ( $v_{0i}$ ) and another for the within-subjects effect of smoking ( $v_{1i}$ ). It is very similar to the model described for EMA data in Moghaddam & Ferguson [15], who examined smoking-related changes in mood. In the proposed extended model, which we refer to as model II, the variances associated with the random subject effects are also modeled in terms of covariates. To allow a subject-level covariate  $w_i$  (e.g. gender or grade) to influence these variances we utilize a log-linear representation, as has been described in the context of heteroscedastic (fixed-effects) regression models [28,29], namely:

$$\sigma_{v_{0i}}^2 = \exp(\alpha_{00} + \alpha_{01}w_i), \quad (\text{eqn 6})$$

$$\sigma_{v_{1i}}^2 = \exp(\alpha_{10} + \alpha_{11}w_i). \quad (\text{eqn 7})$$

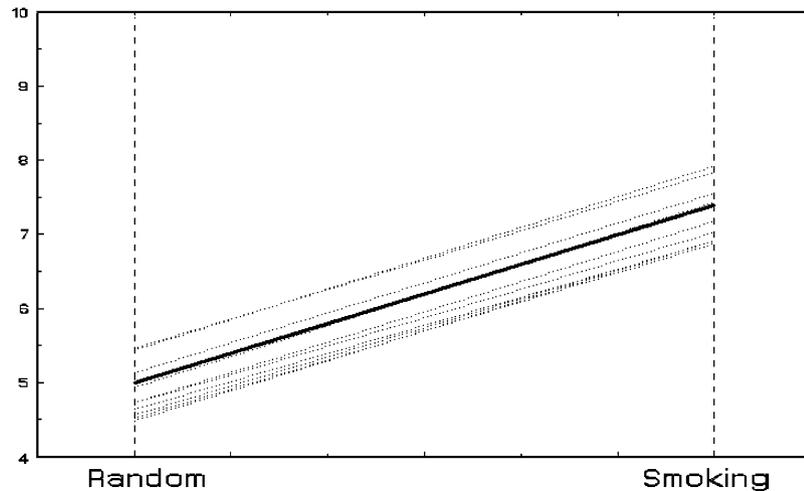
A reason for using this representation is that the exponential function ensures that the variance being modeled will be greater than zero, as it should be. It is merely a convenient mathematical transformation that is used to yield logical results for these variances. Also, the variances are now subscripted by  $i$  to indicate that their values change depending on the value of the subject-level

covariate  $w_i$  and its coefficients. This variable would also be added typically as a covariate in the regression model (i.e. in  $x_{ij}$ ) to account for its effect on the mean of mood. Additionally, an interaction of  $w_i$  by  $\text{SmkEvent}_{ij}$  could be included to allow the covariate effect on the mean of mood to vary between random prompts and smoking events.

In equation 6, the random-prompt variance equals  $\exp\alpha_{00}$  when the subject-level covariate  $w_i$  equals 0, and is increased or decreased as a function of the covariate  $w_i$  and its coefficient  $\alpha_{01}$ . Specifically, if the coefficient  $\alpha_{01} > 0$ , then the random-prompt mood variance increases as  $w_i$  increases (and vice versa if  $\alpha_{01} < 0$ ). The slope variance (i.e. the heterogeneity of the change in mood associated with smoking events) is modeled in the same way in equation 7. That is, this variance equals  $\exp\alpha_{10}$  when the subject-level covariate  $w_i$  equals 0, and is increased or decreased as a function of this covariate and its coefficient  $\alpha_{11}$ .

Figure 2 presents an illustration where there is considerable individual slope variation. If we consider  $w_i$  to be a dichotomous grouping variable ( $= 0$  or  $1$ ) and let Fig. 1 represent the model for  $w_i = 0$  and Fig. 2 represents the model for  $w_i = 1$ , note that the degree of variation around the random prompt average is the same, and so  $\alpha_{01} = 0$ . However, the degree of slope variation is greatly increased in Fig. 2, relative to Fig. 1, and so  $\alpha_{11} > 0$ . Conversely, consider Fig. 3 in which the slope variation is minimal. Here, if Fig. 3 represents the model for  $w_i = 1$ , then  $\alpha_{01} = 0$  (random prompt variance is the same), but  $\alpha_{11} < 0$  (slope variance is decreased).

To summarize, the proposed model has several avenues for examining the potential association of smoking and mood. First,  $\beta_2$  (i.e. the coefficient for  $\overline{\text{SmkEvent}_i}$ ) represents the association of individuals' level of smoking, as indicated by their proportion of smoking events, with their average mood. In other words,



**Figure 3** Random intercept and slope model with decreased slope variance

**Table 1** Positive and negative affect model I estimates, standard errors (SE) and *P*-values.

Parameter	Positive affect			Negative affect		
	Estimate	SE	<i>P</i> <	Estimate	SE	<i>P</i> <
Intercept $\beta_0$	6.713	0.128	0.0001	3.777	0.149	0.0001
SmkEvent <sub>ij</sub> $\beta_1$	0.432	0.073	0.0001	-0.294	0.078	0.0002
SmkEvent <sub>i</sub> $\beta_2$	0.088	0.671	0.90	-1.244	0.780	0.12
Tuesday $\beta_{Tue}$	-0.053	0.064	0.41	-0.041	0.073	0.58
Wednesday $\beta_{Wed}$	-0.152	0.064	0.02	0.106	0.073	0.15
Thursday $\beta_{Thu}$	-0.140	0.064	0.03	0.165	0.073	0.025
Friday $\beta_{Fri}$	0.027	0.064	0.67	-0.025	0.073	0.74
Saturday $\beta_{Sat}$	0.110	0.065	0.09	-0.096	0.074	0.20
Sunday $\beta_{Sun}$	0.005	0.066	0.94	-0.099	0.075	0.19
Intercept variance (ln units) $\alpha_{00}$	0.414	0.098	0.0001	0.752	0.097	0.0001
SmkEvent variance (ln units) $\alpha_{10}$	-0.818	0.269	0.003	-0.903	0.322	0.006
Intercept, SmkEvent covariance $\sigma_{w_0w_1}$	-0.303	0.094	0.002	-0.431	0.121	0.0005
Error variance $\sigma_\epsilon^2$	2.334	0.038	0.0001	3.045	0.049	0.0001

do adolescents who smoke more (less) have lower (higher) general mood? Secondly,  $\beta_1$  (i.e. the coefficient for SmkEvent<sub>ij</sub>), indicates the degree to which a person's mood differs after smoking (compared to random), controlling for their overall level of smoking and its effect on their overall mood. Interacting a subject-level covariate  $w_i$  with SmkEvent<sub>ij</sub> indicates whether the covariate moderates any smoking-associated change in mood levels. Finally, the variance parameter  $\alpha_{11}$  is of particular interest. It indicates the degree to which the covariate  $w_i$  influences smoking-related change in mood variation. For example, it might be that heavier smokers experience greater mood stabilization when smoking than do light smokers. While several studies have examined the effects of covariates on the degree to which the mean of mood changes with smoking [15,30,31], allowing covariates to influence the *variance* in mood, and not just the level of mood, associated with smoking is the unique feature of the proposed model. In this regard, in a recent review

article, Parrott [32] suggested that mood vacillation and its relationship to nicotine dependency was an important topic for future research. The proposed model aims at providing a statistical tool for examining this suggestion.

## RESULTS

Table 1 lists the results for model I applied to positive and negative affect. In addition to the SmkEvent-related variables, indicator variables are included in the analyses to account for day to day mood variation. Monday was selected as the reference day, and so these six day indicators (named Tuesday, Wednesday, . . . , Sunday) represent mood differences relative to Monday.

For the effect of SmkEvent on mood, the results are very consistent for positive and negative affect. Positive mood is increased significantly ( $\hat{\beta}_1 = 0.432$ ,  $P < 0.0001$ ) and negative mood decreased significantly ( $\hat{\beta}_1 = -0.294$ ,  $P < 0.0002$ ) for smoking events, relative to random

**Table 2** Positive and negative affect model II likelihood-ratio tests.

Model	Positive affect			Negative affect		
	Deviance	$\chi^2_{(4)}$	P<	Deviance	$\chi^2_{(4)}$	P<
Model I	30 967			33 129		
Model II with Male	30 966	1	0.91	33 117	12	0.018
Model II with Grade10	30 956	11	0.027	33 127	2	0.74
Model II with NMR	30 932	35	0.0001	33 089	40	0.0001
Model II with SmkLevel	30 956	11	0.027	33 111	18	0.002

$\chi^2_{(4)}$  = likelihood ratio test statistic for comparison to model I. NMR: negative mood regulation.

prompts. These are within-subjects, rather than between-subjects, effects for both outcomes. Controlling for the proportion of smoking events a subject makes, and its effect on mood, subjects' moods are significantly different when they make smoking reports, relative to their random prompts. For the day effects, Thursday is observed to have significantly lower PA and significantly greater NA, relative to Monday ( $\hat{\beta}_{Thu} = -0.140$ ,  $P < 0.03$  and  $\hat{\beta}_{Thu} = 0.165$ ,  $P < 0.025$ , respectively). Wednesday is also significantly lower on PA than Monday ( $\hat{\beta}_{Wed} = -0.152$ ,  $P < 0.02$ ).

In terms of the variance parameters, it is first worth noting that comparing model I to a simpler random-intercepts model (results not shown), which would assume that the SmkEvent variance and covariance parameters equal zero (i.e.  $\alpha_{10} = \sigma_{v_0v_1} = 0$ ), rejects the simpler model in favor of model I (likelihood ratio  $\chi^2_{(2)} = 39$  and  $30$  for PA and NA, respectively,  $P < 0.0001$  for both). Thus, with reference to our first study aim, there is strong evidence that changes in mood associated with smoking vary from subject to subject. Based on the model I estimates, for PA, the between-subjects variance is  $\exp(0.414) = 1.513$  for random prompts and  $\exp(0.414) + \exp(-0.818) + 2(-0.303) = 1.348$  for smoking events. By similar calculations for NA, the between-subjects variance equals 2.121 for random prompts and 1.665 for smoking events. Thus, between-subjects mood variation (i.e. subject heterogeneity), both positive and negative, is reduced under smoking reports, relative to random prompts. Additionally, the covariance parameters are both negative and highly significant ( $\hat{\sigma}_{v_0v_1} = -0.303$ ,  $P < 0.002$  and  $\hat{\sigma}_{v_0v_1} = -0.431$ ,  $P < 0.005$  for PA and NA, respectively), indicating that higher PA and NA values during random prompts are associated with greater reduction in these moods with smoking events. This could be the result of a floor effect of measurement.

Next, we examined the role of covariates using Model II. As mentioned, the covariates we examined were Male, Grade 10, NMR and SmkLevel. Due to the complexity of the model these covariates were examined one at a time, rather than in one large model. We examined the effects

of these covariates on mood as main effects and as moderators of the SmkEvent effect on mood, both in terms of the mean and variance of mood response. Specifically, in terms of the mean of mood, the model was augmented by including each covariate as well as its interaction with SmkEvent (e.g. the terms  $\beta_3$ Male and  $\beta_4$ Male  $\times$  SmkEvent<sub>*j*</sub> were added to equation 3). Similarly, in terms of variance, we added in the covariate effect on the random prompt variance and the SmkEvent-related variance (e.g. Male in place of  $w_i$  in equations 6 and 7). Thus, the difference between models I and II was the inclusion of four parameters for each covariate, and a likelihood ratio test can be used to test the significance of these four additional parameters. Table 2 lists the model deviance (i.e.  $-2 \log$ -likelihood) values and likelihood-ratio test results.

As can be seen, for PA, the significant covariates are Grade 10, NMR and SmkLevel. Similarly, Male, NMR and SmkLevel are significant for NA. Table 3 lists the additional four-parameter estimates of model II for all covariates, although we will interpret the estimates only from models which were deemed significant by the likelihood-ratio test.

For PA, we see that the significant effect of Grade 10 is in terms of the random prompt variance ( $\hat{\alpha}_{01} = -0.418$ ,  $P < 0.03$ ), with 10th graders having reduced PA mood variation, relative to 9th grade students. NMR has a highly significant positive effect on the mean PA response ( $\hat{\beta} = 0.629$ ,  $P < 0.0001$ ), such that subjects with higher negative mood regulation report significantly higher positive affect. SmkLevel significantly decreases the SmkEvent-related variance of positive affect ( $\hat{\alpha}_{11} = -0.337$ ,  $P < 0.02$ ). Thus, increased smoking level is associated with a reduced degree of positive affect change for smoking events, relative to random prompts. In other words, smoking-related positive affect mood response (i.e. the degree of difference in a subject's positive affect between smoking events and random prompts) is significantly less for more frequent smokers.

In terms of NA, Male has a significant effect on both the mean and variance; males are seen to have significantly lower negative affect scores ( $\hat{\beta} = -0.519$ ,  $P < 0.007$ ) which are less variable ( $\hat{\alpha}_{01} = -0.358$ ,

**Table 3** Positive and negative affect model II estimates, standard errors (SE) and *P*-values.

Variable and parameters	Positive affect			Negative affect		
	Estimate	SE	<i>P</i> <	Estimate	SE	<i>P</i> <
Male: gender (0 = female, 1 = male)						
Main effect $\beta_3$	0.087	0.166	0.60	-0.519	0.191	0.007
Interaction $\beta_4$	-0.120	0.145	0.41	0.117	0.157	0.46
Random prompt variance $\alpha_{01}$	-0.020	0.183	0.92	-0.358	0.179	0.05
SmkEvent-related variance $\alpha_{11}$	-0.215	0.498	0.67	0.484	0.563	0.40
Grade 10: grade (0 = 9th, 1 = 10th)						
Main effect $\beta_3$	-0.008	0.168	0.97	0.044	0.196	0.83
Interaction $\beta_4$	0.231	0.140	0.11	-0.181	0.154	0.25
Random prompt variance $\alpha_{01}$	-0.418	0.182	0.03	0.148	0.179	0.42
SmkEvent-related variance $\alpha_{11}$	1.017	0.682	0.14	-0.105	0.582	0.86
NMR: negative mood regulation (1–5)						
Main effect $\beta_3$	0.629	0.110	0.0001	-0.790	0.128	0.0001
Interaction $\beta_4$	-0.109	0.105	0.31	0.054	0.118	0.65
Random prompt variance $\alpha_{01}$	-0.203	0.150	0.18	-0.193	0.133	0.15
SmkEvent-related variance $\alpha_{11}$	-0.300	0.357	0.41	-0.105	0.428	0.81
SmkLevel: 30-day smoking level (0–5)						
Main effect $\beta_3$	-0.025	0.061	0.69	0.144	0.072	0.05
Interaction $\beta_4$	0.068	0.049	0.17	-0.105	0.053	0.05
Random prompt variance $\alpha_{01}$	0.099	0.060	0.11	-0.007	0.062	0.92
SmkEvent-related variance $\alpha_{11}$	-0.337	0.140	0.02	-0.446	0.153	0.004

$P < 0.05$ ). As for positive affect, NMR has a highly significant effect on the mean response, although for NA it is a negative effect ( $\hat{\beta} = -0.790$ ,  $P < 0.0001$ ), such that subjects with higher negative mood regulation have lower negative affect. Finally, SmkLevel has significant mean and variance effects. In terms of the mean, higher smoking level is associated with higher negative affect scores during random prompts ( $\hat{\beta} = 0.144$ ,  $P < 0.05$ ). However, because the interaction is also significant and negative, this effect goes away during smoking events ( $\hat{\beta} = -0.105$ ,  $P < 0.05$ ). Namely, the effect of SmkLevel on negative mood is essentially zero for smoking events ( $= 0.144 - 0.105 = 0.039$ ). In terms of variance, the effect of SmkLevel on negative affect is very similar to its effect on positive affect. That is, increased smoking level is associated with a diminished degree of negative affect change for smoking events, relative to random prompts ( $\hat{\alpha}_{11} = -0.446$ ,  $P < 0.004$ ). Thus, smoking-related mood response (both positive and negative affect) is decreased significantly for more frequent smokers, relative to less frequent smokers.

## DISCUSSION

This paper has illustrated how mixed models for EMA data can be used to model differences in variances, and not just means, across subject-varying covariates. As such, these models can help to identify predictors of

between-subjects variation, and to test psychological hypotheses about these variances. While standard mixed-model software (e.g. SAS PROC MIXED, SPSS MIXED, HLM, MLwiN) can estimate model I easily (albeit not with the natural log scale conversion of the variance parameters that was used here), estimation of the heterogeneous LMM (model II) goes beyond the capabilities of these software programs. However, SAS PROC NLMIXED can be used for this purpose. At the website <http://www.uic.edu/~hedeker/long.html>, we provide sample syntax and output for maximum likelihood estimation of model II, making this class of models accessible to researchers.

Here, we focused on the degree of mood variation between random prompts and smoking events, and whether covariates influenced this variation among adolescent smokers. One of the key concepts in dependence is the development of tolerance, or the diminishing of effects of a substance with continued use. A common experience reported by both adults and adolescents is mood change after smoking a cigarette, and the equally common notion is that these subjective feelings diminish over time as one's experience with smoking increases and tolerance may develop. However, until now, researchers have examined changes in these subjective experiences primarily through paper-and-pencil, retrospective questionnaire reports. Thus, it has been difficult to document adequately whether adolescents experience mood

changes with smoking, and also how symptoms of dependence develop or with what level of smoking experience. Overall, following smoking, adolescents experienced higher positive affect and lower negative affect than they did at random, non-smoking times. However, our analyses also indicated an increased consistency of subjective mood responses as smoking experience increased and a diminishing of mood change. Our data thus provide one of the few ecologically valid examinations of the development of tolerance. Adolescents' self-reports, in real time, of the degree of their subjective response to smoking varied as a function of their smoking level. Importantly, too, these differences were relatively dramatic and were seen for both positive and negative affect. In fact, smoking level was the only covariate that influenced the variance of smoking-related mood response.

Our results also highlight the importance of examining both mean levels of mood and variability; each may convey important information about the development of nicotine dependence in adolescents. Indeed, in our own work (Weinstein *et al.*, in press [33]), we have found that mood variability predicts smoking escalation even after controlling for mean levels of negative mood. Our results also indicated that adolescents who smoked more had higher background, random negative moods. We cannot yet, however, tease apart whether these affective states may reflect in part withdrawal distress, or general contextual or trait-negative affectivity. Our study is one of the first to examine real-time subjective mood responses to smoking among adolescents who are still relatively early in their smoking careers and light or infrequent smokers (less than 7% of the sample smoked more than five cigarettes a day). As such, this study helps to add important information about the relatively early development of symptoms of dependence, a potential development of tolerance to the mood-regulating effects of smoking. These analyses are limited, however, by their cross-sectional nature, and therefore we must take these results as being only suggestive of longitudinal changes as smoking develops.

Our results are also limited by issues related to generalizability in terms of sampling and design characteristics. Our analysis sample was limited to adolescents who had recorded at least one smoking episode during a 7-day assessment week. In addition, we had excluded from this sample adolescents who had reported smoking more than five cigarettes/day at an initial screening survey. Thus, our sample represents relatively light and infrequent smokers. As such, we were unlikely to capture heavier smoking episodes among youth, which might show different patterns of mood variability and mean levels. Nevertheless, our data, with our focus on the earlier stages of smoking, may be important for shedding light on the development of dependence.

More potential applications of this class of models clearly exist in substance abuse and psychological research. For example, many questions of both normal development and the development of psychopathology address the issue of variability or stability in emotional responses to various situations and contexts. Often, an interest is with the variability of responses an individual gives to a variety of stimuli or situations, and not simply with the overall mean level of responsiveness. The models presented here also allow us to examine hypotheses about cross-situational consistency of responses as well.

In order to estimate variances reliably, one needs a fair amount of both within-subjects and between-subjects data. Modern data collection procedures, such as ecological momentary assessments (EMA) and real-time data captures, provide this opportunity. Such designs are in keeping with the 'bursts of measurement' approach described by Nesselroade [34], who called for such an approach in order to assess individual variability. As noted by Nesselroade, such bursts of measurement increase the research burden in several ways; yet they are necessary for studying individual variation, and allow researchers to examine important research questions that were previously unanswerable. Nevertheless, it is not sufficient to have such data to address these questions, one must also have the appropriate statistical tools. This article has aimed at providing tools for this endeavor.

As this is a relatively new modeling technique, certain limitations and cautions should be mentioned. First, our model assumes that the errors and random effects are distributed normally, and it is unclear how robust is this model to violations of normality. Because of the focus on the variance of the dependent variable, the distributional assumptions might be more critical than in more typical models focusing on differences in means. Thus, use of variance stabilizing transformations could be very helpful and should be explored routinely as a preliminary to analysis. Also, it is unclear how well this model would behave for variables where the mean and variance are highly correlated. Again, transformation of the dependent variable could help this situation to some degree. In our analyses we noted the possibility of floor/ceiling effects, and this would seem to be something to keep in mind for rating scale data. Finally, attention should be paid to outliers and influential observations, as these might have undue effects on estimation of the model parameters, especially the variance parameters. Admittedly, this is an emerging area for mixed models [35–37], and work needs to be conducted to generalize these approaches further; however, a careful inspection for outliers in both the dependent and independent variables should accompany any sophisticated statistical modeling.

## Declarations of interest

None.

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## References

1. Verbeke G., Molenberghs G. *Linear Mixed Models for Longitudinal Data*. New York: Springer-Verlag; 2000.
2. Hedeker D., Gibbons R. D. *Longitudinal Data Analysis*. New York: Wiley; 2006.
3. Fleeson W. Moving personality beyond the person-situation debate. *Curr Direct Psychol Sci* 2004; **13**: 83–7.
4. Hertzog C., Nesselroade J. R. Assessing psychological change in adulthood: an overview of methodological issues. *Psychol Aging* 2003; **18**: 639–57.
5. Martin M., Hofer S. M. Intraindividual variability, change, and aging: conceptual and analytical issues. *Gerontology* 2004; **50**: 7–11.
6. Nesselroade J. R. Intraindividual variability and short-term change. *Gerontology* 2004; **50**: 44–7.
7. Stone A., Shiffman S. Ecological momentary assessment (EMA) in behavioral medicine. *Ann Behav Med* 1994; **16**: 199–202.
8. Smyth J. M., Stone A. A. Ecological momentary assessment research in behavioral medicine. *J Happiness Stud* 2003; **4**: 35–52.
9. Scollon C. N., Kim-Prieto C., Diener E. Experience sampling: promises and pitfalls, strengths and weaknesses. *J Happiness Stud* 2003; **4**: 5–34.
10. Feldman Barrett L., Barrett D. An introduction to computerized experience sampling in psychology. *Soc Sci Comp Rev* 2001; **19**: 175–85.
11. Bolger N., Davis A., Rafaeli E. Diary methods: capturing life as it is lived. *Ann Rev Psychol* 2003; **54**: 579–616.
12. Walls T. A., Schafer J. L. *Models for Intensive Longitudinal Data*. New York: Oxford University Press; 2006.
13. Nesselroade J. R., Schmidt McCollam K. M. Putting the process in developmental processes. *Int J Behav Dev* 2000; **24**: 295–300.
14. Schwartz J. E., Stone A. The analysis of real-time momentary data: a practical guide. In: Stone A. A., Shiffman S. S., Atienza A., Nebeling L., editors. *The Science of Real-Time Data Capture: Self-Report in Health Research*. Oxford, England: Oxford University Press; 2007, p. 76–113.
15. Moghaddam N. G., Ferguson E. Smoking, mood regulation, and personality: an event-sampling exploration of potential models and moderation. *J Pers* 2007; **75**: 451–78.
16. Penner L. A., Shiffman S., Paty J. A., Fritzsche B. A. Individual differences in intraperson variability in mood. *J Pers Soc Psychol* 1994; **66**: 712–21.
17. Hedeker D., Berbaum M., Mermelstein R. J. Location-scale models for multilevel ordinal data: between- and within-subjects variance modeling. *J Prob Stat Sci* 2006; **4**: 1–20.
18. Hedeker D., Mermelstein R. J. Mixed-effects regression models with heterogeneous variance: analyzing ecological momentary assessment data of smoking. In: Little T. D., Bovaird J. A., Card N. A., editors. *Modeling Ecological and Contextual Effects in Longitudinal Studies of Human Development*. Mahwah, NJ: Erlbaum; 2007, p. 183–206.
19. Kassel J. D., Stroud L. R., Paronis C. A. Smoking, stress, and negative affect: correlation, causation, and context across stages of smoking. *Psychol Bull* 2003; **129**: 270–304.
20. Khantzian E. J. The self-medication hypothesis of substance use disorders: a reconsideration and recent applications. *Harvard Rev Psychiatry* 1997; **4**: 231–44.
21. Chassin L., Presson C. C., Rose J., Sherman S. J. What is addiction? Age-related differences in the meaning of addiction. *Drug Alcohol Depend* 2007; **87**: 30–8.
22. Kassel J. D., Hankin B. L. Smoking and depression. In: Steptoe A., editor. *Depression and Physical Illness*. Cambridge, England: Cambridge University Press; 2006, p. 321–47.
23. Tiffany S. T., Conklin C. A., Shiffman S., Clayton R. R. What can dependence theories tell us about assessing the emergence of tobacco dependence? *Addiction* 2004; **99**: 78–86.
24. Watson D., Tellegen A. Toward a consensual structure of mood. *Psychol Bull* 1985; **98**: 219–35.
25. Watson D., Wiese D., Vaidya J., Tellegen A. The two general activation systems of affect: structural findings, evolutionary considerations, and psychobiological evidence. *J Pers Soc Psychol* 1999; **76**: 820–38.
26. Catanzaro S. J., Mearns J. Measuring generalized expectancies for negative mood regulation: initial scale development and implications. *J Pers Assess* 1990; **54**: 546–63.
27. Begg M. B., Parides M. K. Separation of individual-level and cluster-level covariate effects in regression analysis of correlated data. *Stat Med* 2003; **22**: 2591–602.
28. Harvey A. C. Estimating regression models with multiplicative heteroscedasticity. *Econometrica* 1976; **44**: 461–5.
29. Aitkin M. Modeling variance heterogeneity in normal regression using GLIM. *Appl Stat* 1987; **36**: 332–9.
30. Adan A., Sanchez-Turet M. Effects of smoking on diurnal variations of subjective activation and mood. *Hum Psychopharmacol Clin Exp* 2000; **15**: 287–93.
31. Adan A., Prat G., Sanchez-Turet M. Effects of nicotine dependence on diurnal variations of subjective activation and mood. *Addiction* 2004; **98**: 1599–607.
32. Parrott A. C. Nicotine psychobiology: how chronic-dose prospective studies can illuminate some of the theoretical issues from acute-dose research. *Psychopharmacology* 2006; **184**: 567–76.
33. Weinstein, S., Mermelstein, R., Shiffman, S., Flay, B. Mood variability and cigarette smoking escalation among adolescents. *Psychol Addict Behav*; in press; 2009.
34. Nesselroade J. R. The warp and woof of the developmental fabric. In: Downs R., Liben L., Palarmo D., editors. *Visions of Development, the Environment, and Aesthetics: The Legacy of Joachim F. Wohlwill*. Hillside, NJ: Earlbaum; 1991, p. 213–40.
35. Langford I. H., Lewis T. Outliers in multilevel data. *J R Stat Soc Ser A* 1998; **161**: 121–60.
36. Shi L., Ojeda M. M. Local influence in multilevel regression for growth curves. *J Multivariate Anal* 2004; **91**: 282–304.
37. Longford N. T. Simulation-based diagnostics in random-coefficient models. *J R Stat Soc Ser A* 2001; **164**: 259–73.